

Applied Stochastic Processes

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1. **Problem:** Find $E(\int_0^T a(t)X(t) dt)^n$, $n = 0, 1, \dots$, where X is a Gaussian process with continuous covariance $R(s, t)$, mean zero, and a is a continuous function.

Solution: Note that $a(t)X(t)$ is also a Gaussian process with continuous covariance $a(s)a(t)R(s, t)$ and mean zero, so without loss of generality we need only consider $M = \int_0^T X(t) dt$.

To find $E(M^n)$, observe that

$$M^n = \lim_{\Delta \rightarrow 0} \sum_{j=1 \dots m} X(i_{j_1}) \dots X(i_{j_n}) \cdot (i_{j_1} - i_{j_1-1}) \dots (i_{j_n} - i_{j_n-1}),$$

where the $\{i_j\}$ are partitions of $[0, T]$, and $\Delta_k = \max(i_{j_k} - i_{j_k-1})$.

Taking expectations,

$$\begin{aligned} E(M^n) &= E \left(\lim_{\Delta \rightarrow 0} \sum_{j=1 \dots m} X(i_{j_1}) \dots X(i_{j_n}) \cdot (i_{j_1} - i_{j_1-1}) \dots (i_{j_n} - i_{j_n-1}) \right) \\ &= \lim_{\Delta \rightarrow 0} E \left(\sum_{j=1 \dots m} X(i_{j_1}) \dots X(i_{j_n}) \cdot (i_{j_1} - i_{j_1-1}) \dots (i_{j_n} - i_{j_n-1}) \right) \\ &= \lim_{\Delta \rightarrow 0} \sum_{j=1 \dots m} R(i_{j_1}, \dots, i_{j_n}) \cdot (i_{j_1} - i_{j_1-1}) \dots (i_{j_n} - i_{j_n-1}) \\ &= \int_{[0, T]^n} R(t_1, \dots, t_n) dt. \end{aligned}$$

This is a Gaussian process, and so the covariance $R(t_1, \dots, t_n)$ can be expressed in terms of the $R(t_j, t_k)$. Doing this and simplifying,

$$\begin{aligned} E(M^2) &= \int_0^T \int_0^T R(s, t) ds dt \\ E(M^{2n+1}) &= 0 \\ E(M^{2n}) &= \frac{1}{n!} \binom{2n}{2, \dots, 2} E(M^2)^n. \end{aligned}$$

In the language of physics and the annihilation and creation operators

$$E(M^n) = \langle 0 | (a + a^*)^n | 0 \rangle \cdot E(M^2)^{n/2},$$

which is the number of 1-valent Feynman diagrams with n vertices multiplied by the volume of the covariance ellipsoid. This is a nice example of a quantum field theory calculation done in a finite-dimensional, rather than infinite-dimensional, Hilbert space.

□

2. **Problem:** Prove that $\sup_{0 < t < 1} (W(t)) > 0$ with probability 1 for a standard Brownian motion W by either using the optional sampling theorem or the reflection principle. Also find $\sup_{\tau} E(W(\tau))$, where the supremum is over all stopping times $\tau \leq 1$.

Solution - Reflection:

Let $a > 0$. By the reflection principle

$$\begin{aligned} p_a &= P\left(\sup_{0 < t < 1} \{W(t)\} \geq a\right) \\ &= 2 \cdot P(W(1) \geq a) \\ &= \frac{2}{\sqrt{2\pi}} \int_a^{\infty} e^{-x^2/2} dx. \end{aligned}$$

It follows that

$$\begin{aligned} P\left(\sup_{0 < t < 1} \{W(t)\} > 0\right) &= \lim_{a \rightarrow 0^+} p_a \\ &= \lim_{a \rightarrow 0^+} \frac{2}{\sqrt{2\pi}} \int_a^{\infty} e^{-x^2/2} dx \\ &= 1. \end{aligned}$$

□

Solution - Optional Stopping Theorem:

Let $\tau_a = \min\{t > 0 : W(t) = a\}$, and let $\tau = \tau_0 \wedge 1$ and $p = P(\tau < 1)$. Now we have that $E(W(\tau)) = p \cdot 0 + (1 - p) \cdot e$, where

$$e = E(W(1) \mid W(t) < 0 \text{ for } 0 < t < 1) \implies e < 0.$$

But by the optional stopping theorem $E(W(\tau)) = E(W(0)) = 1$, and so $p = 1$.

□

3. **Problem:** Show that the simple one-dimensional random walk $S_0 = 0, S_n = \sum_{j=0}^n X_j$, where the X_j are i.i.d. Bernoulli r.v.'s with $P(X_j = \pm 1)$, hits every integer $a > 0$ with probability 1 by using optional sampling and the stopping times $\tau = \min(\tau_{-N}, \tau_a)$, where for any b , τ_b is the first n such that $S_n = b$, or ∞ if there is no such n .

Solution: Let $p = P(\tau = \tau_a)$, then $E(S_\tau) = p \cdot a + (1 - p) \cdot (-N)$. By the optional sampling theorem $E(S_\tau) = E(S_0) = 0$, and so

$$p \cdot a + (1 - p) \cdot (-N) = 0 \implies p = N/(a + N).$$

Letting $N \rightarrow \infty$, we see that $p \rightarrow 1$, and so this random walk hits every $a > 0$ with probability 1. \square

4. **Problem:** Find the expectation of the random variable $\tau = \min(\tau_{-N}, \tau_a)$ in the problem above by making up a martingale that will give you the answer via the optional sampling theorem.

Solution: Consider the martingale $Y_n = S_n^2 - n$ and let p be as above. Here

$$E(Y_\tau) = p \cdot (a^2 - E(\tau_a)) + (1 - p) \cdot (N^2 - E(\tau_{-N})).$$

Again, by the optional sampling theorem $E(Y_\tau) = E(Y_0) = 0$, and so $p \cdot (a^2 - E(\tau_a)) + (1 - p) \cdot (N^2 - E(\tau_{-N})) = 0$, giving

$$E(\tau) = p \cdot E(\tau_a) + (1 - p) \cdot E(\tau_{-N}) = p \cdot a^2 + (1 - p) \cdot N^2.$$

From the previous result $p = N/(a + N)$, and so finally

$$E(\tau) = aN(N + a)/(a + N) = aN.$$

It follows that although this random walk hits every $a > 0$ with probability 1, the expected time to the first hit is infinite.

As a final comment, the martingale $Z_n = S_n^4 - 2n \cdot S_n^2 - n^2 + 2n$ gives the result

$$E(\tau^2) = \frac{1}{3}aN(a^2 - aN + N^2 + 2).$$

\square